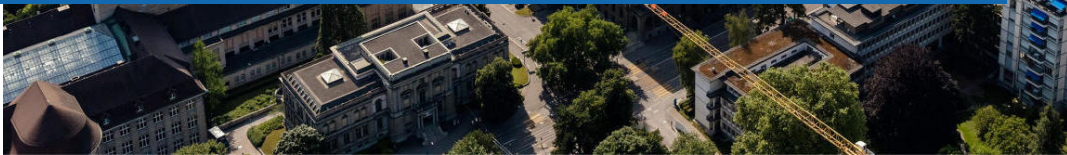




# Stochastic Data-Driven Predictive Control: Regularization, Estimation, and Constraint Tightening

Mingzhou Yin, Andrea Iannelli, Roy S. Smith

July 17, 2024, SYSID 2024



# The LTI stochastic data-driven predictive control problem

- **Data-driven:** Input-disturbance-output trajectory in place of model parameters

$$Z = \begin{bmatrix} z_0^d & \cdots & z_{M-1}^d \end{bmatrix}, \quad z_i^d := \text{col} \left( u_{t_i}^d, \dots, u_{t_i+L-1}^d \mid w_{t_i}^d, \dots, w_{t_i+L-1}^d \mid y_{t_i}^d, \dots, y_{t_i+L-1}^d \right)$$

- **Predictive:**  $L = L_0$  (init. cond. length) +  $L'$  (prediction horizon)

$$\hat{\mathbf{y}}^t \sim \mathcal{D} \left( \underbrace{(\hat{\mathbf{u}}^t := (\hat{u}_k^t)_{k=0}^{L'-1})}_{\text{control inputs}}, \underbrace{\mathbf{w}^t := (w_k)_{k=t-L_0}^{t+L'-1}}_{\text{disturbance sequence}}, \underbrace{\mathbf{u}_{\text{ini}}^t := (u_k)_{k=t-L_0}^{t-1}, \mathbf{y}_{\text{ini}}^t := (y_k)_{k=t-L_0}^{t-1}}_{\text{input \& output init. cond.}} \right)$$

- **Control:** Receding horizon control

$$\min_{(\hat{u}_k^t)_{k=0}^{L'-1}} J_t := \sum_{k=0}^{L'-1} \left\| \hat{u}_k^t \right\|_R^2 + \mathbb{E} \left[ \sum_{k=0}^{L'-1} \left\| \hat{y}_k^t - r_{t+k} \right\|_Q^2 \right]$$

s.t.  $\hat{u}_k^t \in \mathcal{U}_{t+k}, H^{t+k} \hat{y}_k^t \leq q^{t+k}$  with high probability,  $\forall k = 0, \dots, L' - 1$

# The LTI stochastic data-driven predictive control problem

## Stochastic: Two sources of uncertainties

- Zero-mean i.i.d. noise  $\text{cov}(v_t) = \sigma^2 \mathbb{1}_{n_y} \rightarrow$  uncertainties in signal matrix  $Z$  & output initial condition  $\mathbf{y}_{\text{ini}}^t : \mathbb{E}[\mathbf{y}_{\text{ini}}^t] = \bar{\mathbf{y}}_{\text{ini}}^t, \text{cov}(\mathbf{y}_{\text{ini}}^t) = P_t$
- Uncertainties in online disturbance sequence  $\mathbf{w}^t : \mathbb{E}[\mathbf{w}^t] = \bar{\mathbf{w}}^t, \text{cov}(\mathbf{w}^t) = \Sigma_w$

## Agenda:

- Accurate predictor under multiple uncertainties
- Tractable formulation of the expected cost
- Tractable formulation of output constraints

# The stochastic predictor<sup>1</sup>

## Stochastic data-driven prediction

Consider augmented inputs  $\psi_t := \text{col}(u_t, w_t)$ . The distribution of stochastic output prediction can be characterized by  $\mathbb{E}[\hat{\mathbf{y}}^t | g^t] = \bar{\mathbf{y}}^t$ ,  $\text{cov}(\hat{\mathbf{y}}^t | g^t) = \Sigma^t$ , where

$$\bar{\mathbf{y}}^t := Y_f g^t - \Gamma(Y_p g^t - \bar{\mathbf{y}}_{\text{ini}}^t),$$

$$\Sigma^t := \Gamma P_t \Gamma^\top + \Gamma_w \Sigma_w \Gamma_w^\top + \|g^t\|_2^2 T,$$

$$T := \sigma^2 (\Gamma \Gamma^\top + \mathbb{1}_{n_y L'}), \quad \Gamma_w := (Y_f - \Gamma Y_p) R_3,$$

$$\Gamma := \text{col}(CA^{L_0}, \dots, CA^{L_0-1}) \text{col}(C, \dots, CA^{L_0-1})^\dagger.$$

<sup>1</sup>Yin, M., Iannelli, A., and Smith, R.S. (2022). Data-driven prediction with stochastic data: Confidence regions and minimum mean-squared error estimates. In European Control Conference (ECC), 853–858.

# The stochastic predictor

- $g^t$  can be seen as a hyperparameter, typically selected by solving

$$g^t := \operatorname{argmin}_g \left\| Y_p g - \bar{\mathbf{y}}_{\text{ini}}^t \right\|_S^2 + \lambda \|g\|_2^2 = \begin{bmatrix} R_1 & R_2 & R_3 & R_4 \end{bmatrix} \operatorname{col} \left( \mathbf{u}_{\text{ini}}^t, \mathbf{u}^t, \bar{\mathbf{w}}^t, \bar{\mathbf{y}}_{\text{ini}}^t \right)$$

s.t.  $\Psi g = \operatorname{col} \left( \mathbf{u}_{\text{ini}}^t, \hat{\mathbf{u}}^t, \bar{\mathbf{w}}^t \right)$

- The model parameter  $\Gamma$  can be estimated consistently by  $\hat{\Gamma}_Z = Y_f R_4 (Y_p R_4)^{-1}$  under mild conditions

# Stochastic control cost

## Expected control cost

The expected control cost is quadratic with

$$J_t = \left\| \hat{\mathbf{u}}^t \right\|_{\bar{R}}^2 + \left\| \bar{\mathbf{y}}^t - \mathbf{r}^t \right\|_{\bar{Q}}^2 + \text{tr} \left( \bar{Q} T \right) \left\| g^t \right\|_2^2 + \text{constant},$$

where  $\bar{R} := \mathbb{1}_{L'} \otimes R$ ,  $\bar{Q} := \mathbb{1}_{L'} \otimes Q$ ,  $\mathbf{r}^t := (r_{t+k})_{k=0}^{L'-1}$ ,  $T := \sigma^2 \left( \Gamma \Gamma^\top + \mathbb{1}_{n_y L'} \right)$ .

- The  $\|g^t\|_2^2$ -regularization term is commonly seen in DDPC literature, but this result provides a practical approach to selecting the weighting factor

# Initial condition estimation

- In standard DDPC, the output initial condition  $y_{\text{ini}}^t$  is directly measured  
⇒ constant covariance = measurement error
- In MPC, the initial condition  $x_t$  is estimated from both measurement  $y_t$  and previous prediction  $x_{t|t-1}$   
⇒ diminishing error covariance
- **Idea:** Update  $y_{\text{ini}}^t$  with Kalman-filtered measurement from previous prediction
- **Method:** Consider the stochastic predictor as a non-minimal state-space “model” with “state”

$$\bar{x}_t := \text{col} \left( u_{t-L_0}, \dots, u_{t-1}, y_{t-L_0}^0, \dots, y_{t-1}^0 \right)$$

# Initial condition estimation

$$\left\{ \begin{array}{l} \bar{x}_{t+1} = \begin{bmatrix} \Lambda^{n_u} & \mathbf{0} \\ \mathbf{0} & \Lambda^{n_y} \end{bmatrix} \bar{x}_t + \begin{bmatrix} \mathbf{0} \\ \hat{u}_0^t \\ \mathbf{0} \\ \bar{y}_0^t \end{bmatrix} + \begin{bmatrix} \mathbf{0} \\ \mathbf{0} \\ \mathbf{0} \\ e_0^t \end{bmatrix}, \\ \zeta_{t+1} = \begin{bmatrix} \mathbf{0} & \mathbb{1}_{n_y} \end{bmatrix} \bar{x}_{t+1} + v_t = y_t^0 + v_t = y_t \end{array} \right. \quad \left| \quad \bar{x}_t = \begin{bmatrix} u_{t-L_0} \\ \vdots \\ u_{t-1} \\ y_{t-L_0}^0 \\ \vdots \\ y_{t-1}^0 \end{bmatrix}$$

- $\Lambda$ : upper shift operator
- $e_0^t$ : one-step-ahead prediction error with covariance  $\Sigma_0^t$
- $v_t$ : measurement noise with variance  $\sigma^2 \mathbb{1}_{n_y}$
- Standard Kalman filter design can be done



# Chance constraint satisfaction

Both element-wise chance constraints

$$\Pr(h_i^{t+k\top} \hat{y}_k^t \leq q_i^{t+k}) \geq p, \forall i = 1, \dots, n_c, k = 0, \dots, L' - 1$$

and set-wise chance constraints

$$\Pr(H^{t+k} \hat{y}_k^t \leq q^{t+k}) \geq p, \forall k = 0, \dots, L' - 1$$

can be guaranteed by constraint tightening.

- Unlike usual uncertainty assumptions, error depends on inputs via  $g^t$
- Online calculation of constraint tightening is required

# Chance constraint satisfaction

## Constraint tightening

Define augmented linear constraints  $\bar{H}^t \mathbf{y} \leq \bar{q}^t$ . The constraint

$$\bar{q}^t - \bar{H}^t \bar{\mathbf{y}}^t \geq \mu \sqrt{\text{diag} \left( \bar{H}^t \Sigma^t \bar{H}^{t \top} \right)} \quad (\Delta)$$

guarantees the satisfaction of 1) element-wise chance constraints if  $\mu \geq \sqrt{\frac{1}{1-p} - 1}$  and 2) set-wise chance constraints if  $\mu \geq \sqrt{\frac{n_y}{1-p}}$ .

**Proof sketch:** 1) uses one-sided Chebyshev's inequality

$$\Pr \left( \bar{h}_i^t \hat{\mathbf{y}}^t - \bar{h}_i^t \bar{\mathbf{y}}^t \leq \sqrt{\frac{1}{1-p} - 1} \cdot \text{std} \left( \bar{h}_i^t \hat{\mathbf{y}}^t \right) \right) \geq p, \quad \forall i;$$

2) uses multi-dimensional Chebyshev's inequality

$$\Pr \left( \mathbf{e}_k^{t \top} \left( \Sigma_k^t \right)^{-1} \mathbf{e}_k^t \leq \frac{n_y}{1-p} \right) \geq p$$

# Chance constraint satisfaction

- Unfortunately, the tightened constraints are non-convex
- Convex surrogate obtained by using  $\sqrt{\sum_i a_i} \leq \sum_i \sqrt{a_i}$

## Convex surrogate of chance constraints

( $\Delta$ ) is guaranteed by second-order cone constraint

$$\bar{q}^t - \bar{H}^t \bar{y}^t \geq \mu \left( \mathbf{c}_1 + \mathbf{c}_2 \left\| g^t \right\|_2 \right),$$

where

$$\mathbf{c}_1 := \sqrt{\text{diag} \left( \bar{H}^t (\Gamma P_t \Gamma^\top + \Gamma_w \Sigma_w \Gamma_w^\top) \bar{H}^{t^\top} \right)}, \quad \mathbf{c}_2 := \sqrt{\text{diag} \left( \bar{H}^t T \bar{H}^{t^\top} \right)}.$$

# Stochastic indirect DDPC

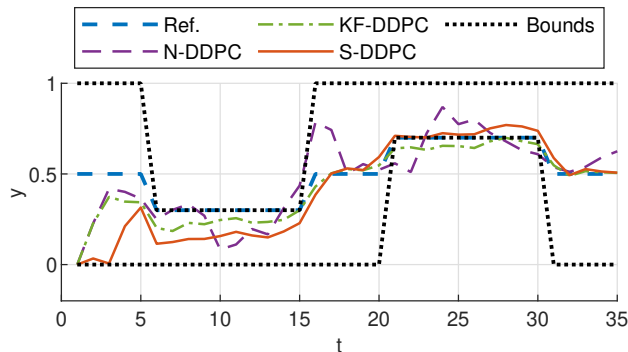
- 1: Select a data-driven predictor and calculate predictor parameters  $R_1, R_2, R_3, R_4, \hat{\Gamma}_Z, T$ .
- 2: Initialize the Kalman filter.
- 3: **for**  $t \leftarrow 0, 1, \dots$  **do**
- 4:     Update Kalman filtered initial conditions  $\text{col}(\mathbf{u}_{\text{ini}}^t, \mathbf{y}_{\text{ini}}^t) \leftarrow \bar{x}_{t,t}, P_t \leftarrow P_{t,t}$ .
- 5:     Solve

$$\begin{aligned} \hat{\mathbf{u}}^t &\leftarrow \underset{\hat{\mathbf{u}}^t}{\text{argmin}} \quad \|\hat{\mathbf{u}}^t\|_{\bar{R}}^2 + \|\bar{\mathbf{y}}^t - \mathbf{r}^t\|_{\bar{Q}}^2 + \text{tr}(\bar{Q}T) \|g^t\|_2^2 \\ \text{s.t.} \quad g^t &= [R_1 \quad R_2 \quad R_3 \quad R_4] \text{col}(\mathbf{u}_{\text{ini}}^t, \mathbf{u}^t, \bar{\mathbf{w}}^t, \bar{\mathbf{y}}_{\text{ini}}^t), \\ \bar{\mathbf{y}}^t &= Y_f g^t - \Gamma(Y_p g^t - \bar{\mathbf{y}}_{\text{ini}}^t), \\ \bar{q}^t - \bar{H}^t \bar{\mathbf{y}}^t &\geq \mu(\mathbf{c}_1 + \mathbf{c}_2 \|g^t\|_2), \\ \hat{u}_k^t &\in \mathcal{U}_{t+k}, \quad \forall k = 0, \dots, L' - 1. \end{aligned}$$

- 6:     Apply  $u_t = \hat{u}_0^t$  to the system and measure  $y_t$ .
- 7: **end for**

# Numerical example

- Fourth-order dynamics
- $L_0 = 4, L' = 10, Q = 20,$   
 $R = 1, \sigma^2 = 0.01, p = 0.95,$   
 $\bar{\mathbf{w}}^t = \mathbf{0}, \Sigma_w = 0.001 \cdot \mathbb{I}$
- Gaussian noise and disturbance
- 500 offline data points
- Lower and upper output bounds
- No input constraints

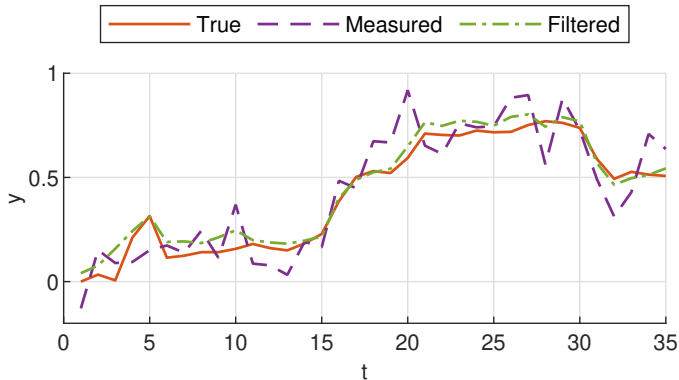


**N-DDPC:** nominal DDPC

**KF-DDPC:** DDPC with Kalman filtering

**S-DDPC:** DDPC with KF & constraint tightening

# Numerical example

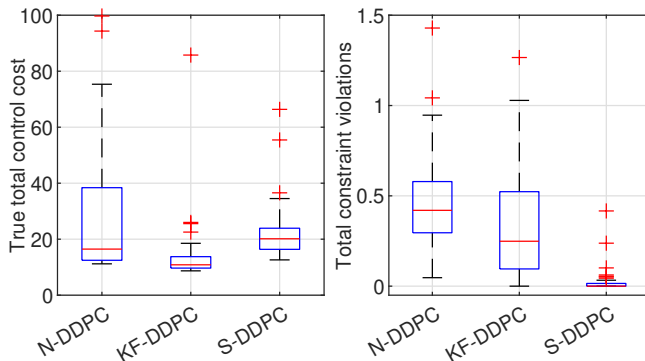


**N-DDPC:** nominal DDPC

**KF-DDPC:** DDPC with Kalman filtering

**S-DDPC:** DDPC with KF & constraint tightening

# Numerical example



- 50 Monte Carlo simulations
- Constraint violation  
$$:= \sum_t \max(H^t y_t - q^t, 0)$$

**N-DDPC:** nominal DDPC

**KF-DDPC:** DDPC with Kalman filtering

**S-DDPC:** DDPC with KF & constraint tightening

## Stochastic Data-Driven Predictive Control: Regularization, Estimation, and Constraint Tightening

- A tuning-free regularizer design in the control cost
- Improved initial condition estimation by Kalman filtering
- Reliable constraint satisfaction by constraint tightening